A Bayesian model for the selection of sample size in clinical trials

Sebastian Jobjörnsson

Chalmers University of Technology

June 5, 2014

(ロ) (同) (三) (三) (三) (○) (○)

Introduction

- IDEAL = Integrated Design and Analysis of small population group trials.
- Overview
 - 1 Basic introduction to Bayesian decision theory.
 - 2 Description of the decision situation.
 - 3 Description of a statistical model.
 - 4 Application of backward induction to the model.
 - 5 Conclusions and discussion of extensions to the model.

(ロ) (同) (三) (三) (三) (○) (○)

BDT = Subjective Bayesian statistics

+ Principle of Maximum Expected Utility

(日) (日) (日) (日) (日) (日) (日)

This can be expanded somewhat:

- Probabilites are degrees of belief.
- Priors on unknown parameters.
- Actions should maximize expected utility.
- Statistical model + Utility function ⇒ Optimal policy.

Decision situation for a clinical trial

- Two treatments: one new, one old.
- Assumption: mean effect for old is 0.
- Assumption: variance for old effect is very small.

Let *N* be the size of the total population.

Let X_k be the effect observed for subject k if given the new treatment.

The agent should do the following:

- 1 Choose a sample size *n* in a clinical trial, $0 \le n \le N$, resulting in observed effects X_1, \ldots, X_n .
- 2 Decide on treatment *d* (1 for new and 0 for old), for N n remaining subjects, resulting in observed effects X_{n+1}, \ldots, X_N or $0, \ldots, 0$.

This is a two-stage decision problem, with *n* and *d* as decision variables.

- **a** X_1, \ldots, X_N i.i.d., with $X_k \mid \mu \sim \mathsf{N}(\mu, \sigma^2)$.
- Assign a conjugate prior distribution μ ~ N(m, s²) to the unknown population mean.

(日) (日) (日) (日) (日) (日) (日)

- The variance σ^2 is assumed known.
- *m* and s^2 are assumed known.

Some notation

Notation for sequences of random variables and realized values of those random variables:

 \blacksquare x_k denotes a realized value of the random variable X_k .

•
$$X^n = (X_k)_{k=1}^n$$
 and $x^n = (x_k)_{k=1}^n$.
• $X^N = (x_k)_{k=n+1}^N$ and $x^N = (x_k)_{k=n+1}^N$.

Notation for sequence means:

$$\overline{X^{n}} = \frac{\sum_{k=1}^{n} X_{k}}{n}$$
$$\overline{x^{n}} = \frac{\sum_{k=1}^{n} x_{k}}{n}$$
$$\overline{X^{N}} = \frac{\sum_{k=n+1}^{N} X_{k}}{N-n}$$
$$\overline{x^{N}} = \frac{\sum_{k=n+1}^{N} x_{k}}{N-n}$$

◆□▶ ◆□▶ ◆目▶ ◆目▶ 目 のへぐ

The utility function for the agent is defined as

$$u(n, \overline{x^{n}}, d, \overline{x^{N}}) = \begin{cases} \sum_{k=1}^{n} x_{k} & \text{if } d = 0\\ \sum_{k=1}^{N} x_{k} & \text{if } d = 1 \end{cases}$$
$$= n * \overline{x^{n}} + d * (N - n) * \overline{x^{N}}$$

 $u(n, \overline{x^n}, d, \overline{x^N})$ is the utility corresponding to

- **1** A choice of sample size *n*, resulting in $X^n = x^n$.
- 2 A choice *d* of treatment for remaining subjects, resulting in $X^N = x^N$ (if d = 1).

(日) (日) (日) (日) (日) (日) (日)

Use backward induction.

For our specific problem, we had:

- 1 Choose a sample size *n* ...
 - ... and observe effects X_1, \ldots, X_n .
- Give the N n remaining subjects the new (d = 1) or old (d = 0) treatment ...

(日) (日) (日) (日) (日) (日) (日)

... and observe effects X_{n+1}, \ldots, X_N (if d = 1).

After all steps, we have $u(n, \overline{x^n}, d, \overline{x^N})$. Now go backwards in time:

• Compute $u(n, \overline{x^n}, d) = \mathbb{E}[u(n, \overline{x^n}, d, \overline{X^N}) \mid n, \overline{x^n}, d].$

Choose optimal *d* by maximizing $u(n, \overline{x^n}, d)$.

• Compute $u(n) = \mathbb{E}[\max_d u(n, \overline{X^n}, d) \mid n]$.

Recall the definition of the utility function as the sum of all observed effects:

$$u(n,\overline{x^n},d,\overline{x^N}) = n * \overline{x^n} + d * (N-n) * \overline{x^N}.$$

This gives

$$u(n, \overline{x^{n}}, d) = \mathbb{E}[n * \overline{x^{n}} + d * (N - n) * \overline{x^{N}} | n, \overline{x^{n}}, d]$$

= $n * \overline{x^{n}} + d * (N - n) * \mathbb{E}[\overline{x^{N}} | n, \overline{x^{n}}, d]$
= $n * \overline{x^{n}} + d * (N - n) * m_{n},$

where $m_n = \mathbb{E}[\overline{X^N} \mid n, \overline{x^n}, d]$ is the updated mean for the posterior distribution of the population mean μ .

Computation of $u(n) = \mathbb{E}[\max_d u(n, \overline{X^n}, d) \mid n]$

From the previous slide:

$$u(n,\overline{x^n},d) = n * \overline{x^n} + d * (N-n) * m_n.$$

This gives:

$$u(n) = \mathbb{E}[\max_{d}(n * \overline{X^{n}} + d * (N - n) * m_{n}) | n]$$

= $\mathbb{E}[n * \overline{X^{n}} + \max_{d}(d * (N - n) * m_{n}) | n]$
= $\mathbb{E}[n * \overline{X^{n}} + (N - n) * \max(0, m_{n}) | n]$
= $\mathbb{E}[n * \overline{X^{n}} | n] + (N - n) * \mathbb{E}[\max(0, m_{n}) | n]$
= $n * m + (N - n) * \mathbb{E}[\max(0, m_{n}) | n].$

Set $t_n^2 = s^2(n/\sigma^2)s_n^2$, where s_n^2 is the updated variance for the population mean μ . Then:

$$u(n) = n * m + (N - n) * \left(t_n^2 * N(0 \mid m, t_n^2) + m * \left(1 - \Phi(0 \mid m, t_n^2) \right) \right).$$

Consider the form of u(n) for the following parameter values:

- Mean *m* of the prior for the population mean μ equals 0.
- Variance s^2 of the prior for the population mean μ equals 1.
- The population variance σ^2 equals 1.

Insertion of these parameter values results in

$$u(n)=\frac{(N-n)}{\sqrt{2\pi}}\sqrt{\frac{n}{n+1}}.$$

< □ > < 同 > < 三 > < 三 > < 三 > < ○ < ○ </p>

Plot of u(n) for example 1

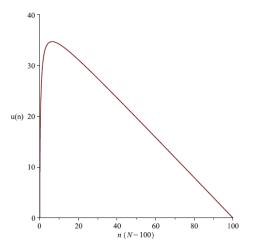


Figure: Plot of u(n) versus n, for N = 100.

(ロ)、

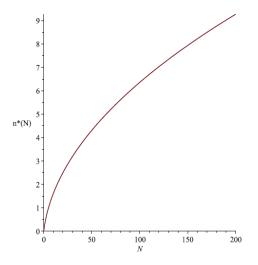
Let $n^*(N)$ be the optimal sample size for a specific N. To find $n^*(N)$, differentiate u(n) w.r.t. n and set the derivative to 0.

This leads to the formula:

$$n^*(N) = -\frac{3}{4} + \frac{1}{4}\sqrt{9+8N}$$

< □ > < 同 > < 三 > < 三 > < 三 > < ○ < ○ </p>

Plot of $n^*(N)$ for example 1



 Introduce a binary regulatory function $r(n, \overline{x^n})$. The agent may choose d = 1 if and only if $r(n, \overline{x^n}) = 1$. Assume the form:

$$r(n,\overline{x^n}) = \begin{cases} 1 & \text{if } \overline{x^n} > \frac{\sigma z_{\alpha}}{\sqrt{n}} \\ 0 & \text{if } \overline{x^n} \le \frac{\sigma z_{\alpha}}{\sqrt{n}} \end{cases}$$

where $\mathbb{P}(Z \ge z_{\alpha}) = \alpha$ if $Z \sim N(0, 1)$. This corresponds to testing $H_0 : \mu \le 0$ against $H_1 : \mu > 0$, at significance level α .

 $u_r(n)$ may be derived as in the case without any requirement.

(日) (日) (日) (日) (日) (日) (日)

Compute u(n) and $u_r(n)$ for the same parameter values as in example 1:

 $m = 0, \sigma^2 = s^2 = 1$. Insertion into the general formulas yields

$$u(n) = \frac{(N-n)}{\sqrt{2\pi}} \sqrt{\frac{n}{n+1}}$$
$$u_r(n) = u(n)e^{-\frac{1}{2}(\frac{z_{\alpha}^2}{n+1})}.$$

Let

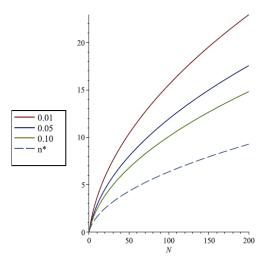
• $n^*(N)$ = Optimal sample size without *r*.

• $n_r^*(N)$ = Optimal sample size with *r*.

The next slide shows a plot of these functions for three different significance levels: $\alpha = 0.01$, $\alpha = 0.05$ and $\alpha = 0.10$.

(日) (日) (日) (日) (日) (日) (日)

Example 2: Plot of optimal sample sizes



▲□▶ ▲□▶ ▲ 三▶ ▲ 三▶ - 三 - のへぐ

Let

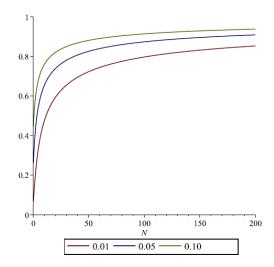
•
$$u^*(N) = u(n^*(N))$$
 = Optimal utility without *r*.

▲□▶ ▲□▶ ▲ 三▶ ▲ 三▶ - 三 - のへぐ

•
$$u_r^*(N) = u_r(n^*(N))$$
 = Optimal utility with *r*.

In order to compare the two cases, we plot the ratio $\frac{u_t^*(N)}{u^*(N)}$ against *N*.

Example 2: Plot of $u_r^*(N)/u^*(N)$ versus N



Conclusions

For our simple model:

- Significant impact of regulatory requirements for small population sizes N.
- 2 The precise loss in expected utility may be computed.

◆□▶ ◆□▶ ▲□▶ ▲□▶ □ のQ@

But the model is too simple!

For the statistical model:

- Drop the assumption of known σ^2 (and use prior).
- Generalize to other distributions for the effect.
- For the utility function:
 - Include cost of trial in the utility function.
 - Change the utility function entirely? (For example, if the agent is a company).
- For the structure of the decision process:
 - More stages, there could be several clinical trials for the same treatment.
 - More agents? At least two seems appropriate: Sponsor and Regulatory Agency.

(ロ) (同) (三) (三) (三) (○) (○)

Go from constant N to random process N_t ?



Thanks for listening!

