



Assessment of randomization procedures based on single sequences under selection and chronological bias

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- 21 out of 63 Orphan drug legislations involve open label studies (Joppi, 2013)
- *treatment comparisons should involve consideration of the potential contribution of bias to the p -value (ICH E9)*
- no recommendation to give scientific arguments for selection of randomization procedure
- no uniform performance of randomization procedures
- Clinical Scenario Evaluation (Benda 2010)





Propose a tool for assessing the impact of selection bias as well as chronological bias on the type one error probability for a given randomization sequence (procedure) and thus enabling a scientific discussion of the appropriate choice of the randomization procedure





Two arm parallel group design

$$y_i = \mu_A Z_i + \mu_B(1 - Z_i) + \tau_i + \epsilon_i, \quad 1 \leq i \leq n_A + n_B \quad (1)$$

- allocation

$$Z_i = \begin{cases} 1 & \text{if patient } i \text{ is allocated to group } A \\ 0 & \text{if patient } i \text{ is allocated to group } B \end{cases}$$

- μ_j expected response under treatment $j = A, B$
- τ_i denotes the fixed unobserved "bias" effect acting on the response of patient i
- errors ϵ_i iid $N(0, \sigma^2)$





Aim: test the hypotheses $H_0 : \mu_A = \mu_B$ vs. $H_1 : \mu_A \neq \mu_B$

use t-Test (under misspecification)

$$T = \frac{\sqrt{\frac{n_A n_B}{n_A + n_B}} (\tilde{y}_A - \tilde{y}_B)}{\frac{1}{n_A + n_B - 2} \left(\sum_{i=1}^n Z_i (y_i - \tilde{y}_A)^2 + \sum_{i=1}^n (1 - Z_i) (y_i - \tilde{y}_B)^2 \right)}$$

where $\tilde{y}_A = \frac{1}{n_A} \sum_{i=1}^n y_i Z_i$; $\tilde{y}_B = \frac{1}{n_B} \sum_{i=1}^n y_i (1 - Z_i)$; $n = n_A + n_B$





Theorem: Under $H_0 : \mu_A = \mu_B$ the type 1 error probability in (1) (under misspecification) for the allocation sequence $\mathbf{Z} = (Z_1, \dots, Z_n)$ is

$$P(|T| > t_{n_A+n_B-2}(1-\alpha/2) | \mathbf{Z}) \\ = F_{n_A+n_B-2, \theta_1, \theta_2}(t_{n_A+n_B-2}(\alpha/2)) + F_{n_A+n_B-2, -\theta_1, \theta_2}(t_{n_A+n_B-2}(\alpha/2)).$$

$F_{n_A+n_B-2, \theta_1, \theta_2}$ denotes the distribution function of the doubly non-central t-distribution with $n_A + n_B - 2$ degrees of freedom and parameters

$$\theta_1 = \frac{1}{\sigma} \sqrt{\frac{n_A n_B}{n_A + n_B}} (\tilde{\tau}_A - \tilde{\tau}_B) \quad \theta_2 = \frac{1}{\sigma^2} \left[\sum_{i=1}^n \tau_i^2 - n_A \tilde{\tau}_A^2 - n_B \tilde{\tau}_B^2 \right]$$

where $\tilde{\tau}_A = \frac{1}{n_A} \sum_{i=1}^n \tau_i Z_i$; $\tilde{\tau}_B = \frac{1}{n_B} \sum_{i=1}^n \tau_i (1 - Z_i)$





Sketch of the proof:

for the given allocation vector $\mathbf{Z} = (Z_1, \dots, Z_n)$

$$\tilde{y}_A - \tilde{y}_B \sim N \left(\mu_A - \mu_B + \tilde{\tau}_A - \tilde{\tau}_B, \sigma^2 \frac{n_A + n_B}{n_A n_B} \right)$$

$$\text{where } \tilde{\tau}_A = \frac{1}{n_A} \sum_{i=1}^n \tau_i Z_i ; \quad \tilde{\tau}_B = \frac{1}{n_B} \sum_{i=1}^n \tau_i (1 - Z_i)$$





for the allocation vector $\mathbf{Z} = (Z_1, \dots, Z_n)$

$$\sum_{i=1}^n Z_i (y_i - \tilde{y}_A)^2 + \sum_{i=1}^n (1 - Z_i) (y_i - \tilde{y}_B)^2 \sim \chi_{n_A + n_B - 2}(\theta_2)$$

with non-centrality parameter

$$\begin{aligned} \theta_2 &= \frac{1}{\sigma} \left(\sum_{i=1}^n Z_i (\tau_i - \tilde{\tau}_A)^2 + \sum_{i=1}^n (1 - Z_i) (\tau_i - \tilde{\tau}_B)^2 \right) \\ &= \frac{1}{\sigma} \left[\sum_{i=1}^n \tau_i^2 - n_A \tilde{\tau}_A^2 - n_B \tilde{\tau}_B^2 \right] \end{aligned}$$





Thus T follows a doubly non-central t distribution with $n_A + n_B - 2$ degrees of freedom and non-centrality parameters (Johnson, Kotz, Balakrishnan, 1995, Robins, 1948)

$$\theta_1 = \frac{1}{\sigma} \sqrt{\frac{n_A n_B}{n_A + n_B}} (\mu_A - \mu_B + \tilde{\tau}_A - \tilde{\tau}_B) = \frac{1}{\sigma} \sqrt{\frac{n_A n_B}{n_A + n_B}} (\tilde{\tau}_A - \tilde{\tau}_B)$$

$$\theta_2 = \frac{1}{\sigma^2} \left[\sum_{i=1}^n \tau_i^2 - n_A \tilde{\tau}_A^2 - n_B \tilde{\tau}_B^2 \right]$$

using the properties of the distribution (Kocherlakota, 1991)

$$F_{\nu, \theta_1, \theta_2}(t) = 1 - F_{\nu, -\theta_1, \theta_2}(-t)$$





- CR** Complete randomization is accomplished by tossing a fair coin, so the probability that patient i will receive treatment 1 is always $\frac{1}{2}$
- RAR** Random Allocation rule, fix total sample size N . Randomize so that half the patients receive treatment 1
- PBR** (Permuted Block Randomization) Implementation of RAR within k Blocks of size $m_s, 1 \leq s \leq k$
- BSD(a)** (Big Stick design) CR allow for imbalance within a limit a





With $N_j(i-1)$ is the number of treatment j assignments after $(i-1)$ assignments and

$$p_A(i-1) = \frac{n_A - N_A(i-1)}{(n_A + n_B) - (N_A(i-1) + N_B(i-1))}, \quad n_A = n_B$$

the selection biasing policy (Tamm, 2012), according to convergence strategy (Blackwell, 1957), $q \in [\frac{1}{2}, 1]$

$$E(y_i) = \begin{cases} \mu_A Z_i + \mu_B(1 - Z_i) + \eta & \text{if } p_A(i-1) > q \\ \mu_A Z_i + \mu_B(1 - Z_i) & \text{if } 1 - q \leq p_A(i-1) \leq q \\ \mu_A Z_i + \mu_B(1 - Z_i) - \eta & \text{if } p_A(i-1) < 1 - q \end{cases}$$

Then

$$\tau_i = \eta [\mathbb{1}_{(q,1]}(p_A(i-1)) - \mathbb{1}_{[0,1-q)}(p_A(i-1))]$$





Table: Empirical type 1 error probability by different cut-off values for a two sided t-test with $2n = 40$; $\alpha = 0.05$, $\beta = 0.2$ and selection effect $\eta = \frac{\delta}{2} = 0.45$ under PBR

Cutoff q	Type 1 Error rate		
	PBR(4)	PBR(8)	PBR(10)
1/2	0.191	0.134	0.118
2/3	0.145	0.134	0.086
1	0.050	0.050	0.048

using SAS with 10 000 replications





Table: Empirical type 1 error probability of a two sided t-test with $2n$; $\alpha = 0.05$, $\beta = 0.2$ and selection bias effect $\eta = \frac{\delta(2n)}{2}$

$2n$	$\delta(2n)$	CR	RAR	PBR(4)	BSD (2)
8	2.381	0.058	0.102	0.141	0.064
20	1.325	0.054	0.082	0.177	0.075
32	1.024	0.055	0.072	0.188	0.083
40	0.909	0.053	0.071	0.195	0.088

using R with 100 000 replications





Tamm (2012) found:

- empirical type I error increases with smaller blocksize under PBR
- empirical type I error decreases with higher q
- elevation are substantial even for $q = 2/3$





- long recruitment time in Rare Diseases, (EMA, 2006)
 - ▶ changes in population characteristics
 - ▶ learning effect in therapy / surgical experience (Hopper, 2007)
 - ▶ change in diagnosis (FDA, 2011), etc.
- special form of accidental bias, when considering a time-heterogeneous covariate

$$\tau_i = \lambda \begin{cases} \frac{i}{n_A+n_B} & \text{linear time trend} \\ \mathbb{1}_{i \geq c}(i) & \text{stepwise trend} \\ \log\left(\frac{i}{n_A+n_B}\right) & \text{log trend} \end{cases}$$





Tamm (2014) found, that

- using t-test in presence of time trends results in conservative test decisions under permuted block randomizations even for small time trends; Blocked ANOVA should be used
- medium block sizes already achieve a good reduction of the inflated type I error rate in worst-case scenarios



Assessment - Selection and chronological bias for special sequences



all Sequences assume $n_A = n_B = n/2$

Sequence	Bias	θ_1	θ_2	type 1 error probability
Alternating 10101...	Selection	$-\frac{\eta}{\sigma}\sqrt{n}/2$	0	0.074
	Time Trend	$\frac{\lambda}{\sigma}\sqrt{n}/2$	$\frac{\lambda^2}{\sigma^2}n(n^2 - 4)/12$	0.000
Separating 111...000...	Selection	$-\frac{\eta}{\sigma}\sqrt{1/n}$	$\frac{\eta^2}{\sigma^2}(n - 2)/n$	0.050
	Time Trend	$\frac{\lambda}{\sigma}\sqrt{n^3}/4$	$\frac{\lambda^2}{\sigma^2}n(n^2 - 4)/48$	0.999
avoid TT 1001 1001..	Selection	$-\frac{\eta}{\sigma}\sqrt{n}/2$	$\frac{\eta^2}{\sigma^2}$	0.050
	Time Trend	0	$\frac{\lambda^2}{\sigma^2}n(-1 + n^2)/12$	0.000

type 1 error probability setting: $n_A = n_B = 6, \eta = \frac{\delta(12)}{4} = 0.45, \lambda = \sigma = 1$





weighted additive (selection and chronological) bias model

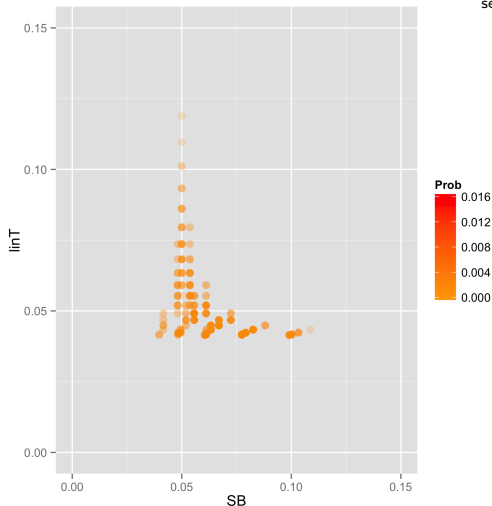
$$\tau_i = \underbrace{\lambda \frac{i}{n_A + n_B}}_{\text{time trend}} + \underbrace{\eta \text{sign}(N_A(i-1) - N_B(i-1))}_{\text{selection bias with } q = \frac{1}{2}}$$

- weights via definition of λ and η
- different shape of time trend can be incorporated
- relaxed version of selection bias possible
- multiplicative could also be done





setting: $n_A = n_B = 6$, $\eta = \frac{\delta(12)}{4} = 0.45$, $\lambda = \sigma = 1$



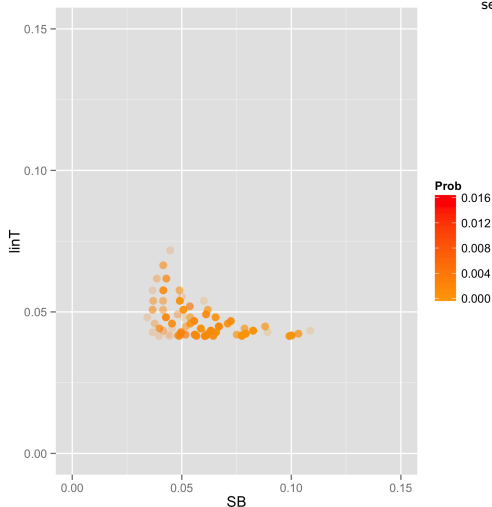
Measure	p value	
	Selection Bias	Linear Time Trend
min	0.0397	0.0415
x05	0.0481	0.0414
x25	0.0518	0.0423
x50	0.0611	0.0449
mean	0.0635	0.0500
x75	0.0724	0.0520
x95	0.0991	0.0736
max	0.1086	0.1188
sd	0.0151	0.0115



Assessment - Assessment of Bias for BSD(2) (N=12)



setting: $n_A = n_B = 6$, $\eta = \frac{\delta(12)}{4} = 0.45$, $\lambda = \sigma = 1$



Measure	p value	
	Selection Bias	Linear Time Trend
min	0.0342	0.0415
x05	0.0370	0.0415
x25	0.0483	0.0420
x50	0.0537	0.0441
mean	0.0557	0.0462
x75	0.0619	0.0481
x95	0.0792	0.0576
max	0.1086	0.0717
sd	0.0129	0.0059



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Figure: RAR

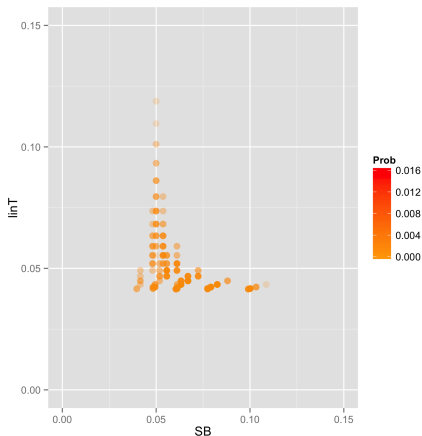
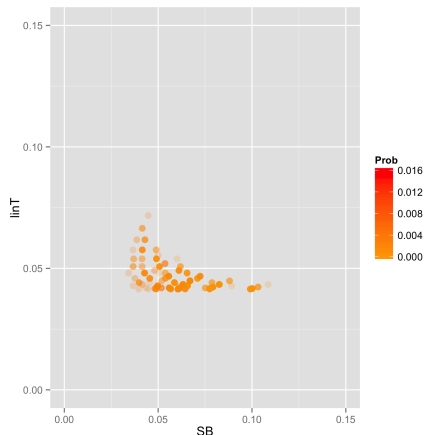


Figure: BSD(2)



setting: $n_A = n_B = 6$, $\eta = \frac{\delta(12)}{4} = 0.45$, $\lambda = \sigma = 1$














- presented a framework for scientific evaluation of randomization procedures on bias
- other randomization procedures are easy to implement
- this evaluation should be part of the TSAP
- dichotomous endpoint / survival endpoint could be evaluated by simulation (*work under progress by Marcia Rückbeil*)
- Cave: this is the wrong test (!) but mirrors the practical situation, use randomization tests to reflect the randomization argument
- R package (*randomizeR*) coming soon, see website (<http://www.ideal.rwth-aachen.de>)





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